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DOI [https://doi.org/10.24144/2616-7700.2022.41\(2\).61-68](https://doi.org/10.24144/2616-7700.2022.41(2).61-68)**T. B. Lysetskyi**

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ORCID: <https://orcid.org/0000-0001-5993-1887>**ON A NUMBER OF EMIGRANTS IN DECOMPOSABLE  
AGE-DEPENDENT BRANCHING PROCESSES**

Decomposable branching process can be viewed as stochastic model for the population with  $N$  types of individuals, splitted into several groups  $G_1, G_2, \dots, G_n$ ,  $n \leq N$ , where each group occupies its own island. Individual of group  $G_i$  may immediately after birth emigrate to island, occupied by group with higher index or stay on the same island. In given paper we consider case with two groups  $G_1$  and  $G_2$ . Each individual has random life duration and distribution of its progeny depends on its age.

We establish asymptotic behaviour of processes that count number of emigrated individuals, depending on criticality of branching subprocess, generated by group  $G_1$ .

**Keywords:** branching process, stochastic additive functional, critical branching process, Perron root, moments.

**1. Introduction.** We will provide short description of decomposable multitype age-dependent branching process with variable transition probabilities. Review of multitype age-dependent processes with variable transition probabilities can be found in [1], chapter 8, while decomposable branching processes were studied, for example, in [2], [3] and [4]. Description of probability space could be given analogically to [5], chapter 6.

Consider a population consisting of  $n$  types  $T_1, T_2, \dots, T_n$ . Each  $T_i$  –  $th$  type particle has random life duration  $\tau_i$  with distribution function

$$P(\tau_i \leq t) = G^i(t), G^i(0+) = 0$$

We will assume that  $G^i(t)$  are absolutely continuous.

Types of particles can be divided into two groups:  $C_1$  which includes particles of types  $T_1, T_2, \dots, T_r$  and  $C_2$  which includes particles (individuals) of types  $T_{r+1}, T_{r+2}, \dots, T_n$ ,  $1 \leq r < n$ . Particles from  $C_2$  at the end of their lives transform into any number of particles from their own group, while particles from  $C_1$  transform into any number of particles of any type. Direct  $C_2$  types descendants of particles from  $C_1$  group types we will call 'emigrants'. Establishing asymptotic behaviour of processes, which describe number of emigrants is the goal of this paper.

Conditional probability (if transformation took place when the age attained by the original particle was  $u$ )  $p_\alpha^i(u)$  of transformation into a set consisting of  $\alpha_i$   $T_i$  –  $th$  type particles,  $i = \overline{1, n}$ , where  $\alpha = (\alpha_1, \alpha_2, \dots, \alpha_n)$  is  $n$  - dimensional vector where the components are non-negative integers. Evolution of particle is defined by joint distribution of random variable  $\tau_i$  and random vector  $v_i = (v_i^1, \dots, v_i^n)$ , which characterise progeny of this particle.

$$P(\tau_i \in B, v_i = \alpha) = \int_B p_\alpha^i(u) dG^i(u).$$

Vector  $\mu_i(t) = (\mu_i^1(t), \dots, \mu_i^n(t))$  denotes number of particles of types  $T_1, T_2, \dots, T_n$  at the moment  $t$ , under the condition that at initial moment there existed one  $T_i - th$  type particle. And let's also assume, that vectors  $\mu^i(t)$  are right continuous.

Let's denote by  $P^i(*)$  and  $E^i(*)$  denote conditional probability and conditional expectation respectively, under condition that at initial moment of time there existed one particle of type  $T_i$ .

Let's also introduce generating functions

$$h^i(t, s) = \sum_{\alpha} P_{\alpha}^i(t) s^{\alpha}$$

and

$$F^i(t, s) = \sum_{\alpha} P^i(\mu_i(t) = \alpha) s^{\alpha}, i = \overline{1, n},$$

$$s = (s_1, \dots, s_n), s^{\alpha} = s_1^{\alpha_1} \cdot \dots \cdot s_n^{\alpha_n},$$

$$F(t, s) = (F^1(t, s), \dots, F^n(t, s)), h(t, s) = (h^1(t, s), \dots, h^n(t, s)).$$

Generating functions  $F^i(t, s)$  satisfy with  $s \leq 1$  (componentwise) and  $t \geq 0$  next system of integral equations [1]

$$F^i(t, s) = s_i(1 - G^i(t)) + \int_0^t h^i(u, F(t - u, s)) dG^i(u), i = \overline{1, n}. \quad (1)$$

**2. Preliminaries.** Define

$$a_j^i(u) = \frac{\partial h^i(u, s)}{\partial s_j} \Big|_{s=1}, A_j^i = \int_0^{\infty} a_j^i(u) dG^i(u),$$

$$b_{jk}^i(u) = \frac{\partial h^i(u, s)}{\partial s_j \partial s_k} \Big|_{s=1}, B_{jk}^i = \int_0^{\infty} b_{jk}^i(u) dG^i(u).$$

We will call described above branching process (b.p.)  $\xi$ . We assume that matrix  $A = \|A_j^i\|_{i,j=\overline{1,n}}$  of moments has a form  $\begin{bmatrix} A_1 & A_{12} \\ \mathbf{0} & A_2 \end{bmatrix}$ , where matrices  $A_1 = \|A_j^i\|_{i,j=\overline{1,r}}$  and  $A_2 = \|A_j^i\|_{i,j=\overline{r+1,n}}$  are irreducible,  $A_{12} = \|A_j^i\|_{i=\overline{1,r}, j=\overline{r+1,n}}$ ,  $\mathbf{0} = (n-r) \times r$  zero matrix.

Since our goal is to describe number of emigrants, we will mainly focus on matrix  $A_1$ .

Let

$$A_{\rho k}^l = \int_0^{\infty} u e^{-\rho u} a_k^l(u) dG^l(u), B_{\rho j k}^l = \int_0^{\infty} e^{-2\rho u} b_{jk}^l(u) dG^l(u),$$

$$M^k = \int_0^{\infty} u dG^k(u), M_a^{lk} = \int_0^{\infty} u a_k^l(u) dG^l(u),$$

$$B = \sum_{l,k,m=1}^r B_{mk}^l v^l u^k u^m, M_a = \sum_{l,k=1}^r M_a^{lk} v^l u^k, M_j = \sum_{l=1}^r v^l \int_0^{\infty} a_j^l(u) dG^l(u),$$

where  $\rho$  denotes such number, that perron root of matrix  $\| \int_0^{\infty} e^{-\rho u} a_k^l(u) dG^l(u) \|_{i,j=\overline{1,r}}$  equals to one,  $u_{\rho} = (u_{\rho}^1, u_{\rho}^2, \dots, u_{\rho}^r)$  and  $v_{\rho} = (v_{\rho}^1, v_{\rho}^2, \dots, v_{\rho}^r)$  denote right and left

eigenvectors of this matrix,  $(v_\rho, u_\rho) = (v_\rho, \mathbf{1}) = \sum_{k=1}^r v_\rho^k = 1$  (in critical case  $\rho = 0$  and we denote  $u_\rho = u, v_\rho = v$ ).

To simplify further investigation we will introduce process  $\xi^t$ , which consists only of particles from group  $C_1$ . By  $\nu_i(t) = (\nu_i^{r+1}(t), \dots, \nu_i^n(t))$  we will denote number of particles of types from group  $C_2$  (emigrants), produced by  $C_1$  type particles by  $t$ . Processes  $\nu_i(t), i = \overline{1, r}$  can be seen as 'product' of  $C_1$  type particles or multidimensional additive functional from branching process  $\xi^t$  (see [7]).

Consider vector  $\tilde{\mu}_i(t) = (\mu_i^1(t), \dots, \mu_i^r(t), \nu_i^{r+1}(t), \dots, \nu_i^n(t)), i = \overline{1, n}$ .

By  $\tilde{F}^i(t, s)$  we will denote generating functions of this vectors. It is easy to see that  $F^j(t, s) = s_j$  for  $j = \overline{r+1, n}$  and formula (1) is valid for  $\tilde{F}^i(t, s), i = \overline{1, r}$  with  $\tilde{F}(t, s) = (\tilde{F}^1(t, s), \dots, \tilde{F}^n(t, s))$ .

According to [1], we have next formulas for moments

$$A_j^i(t) = \frac{\partial F^i(t, s)}{\partial s_j} = \frac{\partial \tilde{F}^i(t, s)}{\partial s_j},$$

$$B_{jk}^i(t) = \frac{\partial^2 F^i(t, s)}{\partial s_j \partial s_k} = \frac{\partial^2 \tilde{F}^i(t, s)}{\partial s_j \partial s_k},$$

$$A_j^i(t) = \delta_j^i(1 - G^i(t)) + \sum_{k=1}^r \int_0^t A_j^k(t - u) a_k^i(u) dG^i(u),$$

$$B_{jk}^i(t) = \sum_{l=1}^r \int_0^t B_{jk}^l(t - u) a_l^i(u) dG^i(u) + \sum_{m,l=1}^r \int_0^t A_j^l(t - u) A_k^m(t - u) b_{lm}^i(u) dG^i(u),$$

for  $i, j, k = \overline{1, r}$ .

For  $A_l^i(t), B_{jl}^i(t)$  and  $B_{il}^i(t)$ , were  $i, j = \overline{1, r}, l = \overline{r+1, n}$  we will have

$$A_l^i(t) = \int_0^t a_l^i(u) dG^i(u) + \sum_{k=1}^r \int_0^t A_l^k(t - u) a_k^i(u) dG^i(u), \tag{2}$$

$$\begin{aligned} B_{jl}^i(t) &= \sum_{m=1}^r \int_0^t B_{jl}^m(t - u) a_m^i(u) dG^i(u) + \\ &+ \sum_{m,d=1}^r \int_0^t A_j^m(t - u) A_l^d(t - u) b_{mq}^i(u) dG^i(u), \end{aligned} \tag{3}$$

$$\begin{aligned} B_{il}^i(t) &= \sum_{m=1}^r \int_0^t B_{il}^m(t - u) a_m^i(u) dG^i(u) + \\ &+ \sum_{m,q=1}^r \int_0^t A_i^m(t - u) A_l^q(t - u) b_{mq}^i(u) dG^i(u). \end{aligned} \tag{4}$$

Next theorem is crucial for establishing asymptotic behaviour of moments. Proof of this theorem can be found in [6].

**Theorem 1.** Let  $M(dy) = \|m_{ij}(dy)\|_{i,j=\overline{1,r}}$  be  $r \times r$  square matrix, components of which are finite non-negative measures on  $[0, +\infty)$  and let vector function  $g(x) = (g_1(x), \dots, g_r(x))$  be such that for some  $\gamma \geq 0$  holds  $\sup_{x \in \mathbb{R}} \frac{g(x)}{\max\{1, x^\gamma\}} < \infty$  and  $\frac{g(x)}{x^\gamma} \xrightarrow{x \rightarrow +\infty} c = (c_1, \dots, c_r)$ . If Perron root of  $M[0, +\infty)$  equals to 1 and  $\int_0^\infty u m_{ij}(du) < \infty$ , then

$$\frac{1}{x^{1+\gamma}} \int_0^x g(x-y) dH(y) \xrightarrow{x \rightarrow +\infty} \frac{c}{(1+\gamma)a} \|u^i v^j\|_{i,j=\overline{1,r}},$$

where  $H(y)$  is renewal matrix, which corresponds to matrix  $M(dy)$ ,  $u$  and  $v$  are right and left eigenvectors of  $M[0, +\infty)$ ,  $a = \left( v, \int_0^\infty y M(dy) u \right)$ .

Applying **Theorem 1** to (2)–(4), one can show that, in critical case, if  $M^j$ ,  $M_a^{jk}$ ,  $B_{lk}^j$  are finite, then

$$A_l^i(t) \sim \frac{u^i M_l t}{M_a}, B_{jl}^i(t) \sim B \frac{u^i v^j M^j M_l t^2}{2(M_a)^2}, B_u^i(t) \sim B \frac{u^i (M_l)^2 t^3}{3(M_a)^3}. \quad (5)$$

We will also need next lemma to prove **Theorem 2**.

**Lemma 1.** If for random variables  $X(t)$  and  $Y(t)$  following conditions are satisfied:

- a)  $X(t) \xrightarrow{t \rightarrow +\infty} X$  in distribution;
- b)  $\lim_{t \rightarrow \infty} (X(t) - Y(t))^2$ ,

then  $Y(t) \xrightarrow{t \rightarrow +\infty} X$  in distribution.

**Proof.** Indeed, from condition b) we get that  $Y(t) = X(t) + \theta(t)$ , where  $\theta(t)$  tends to zero in square mean, therefore in distribution, so

$$\lim_{t \rightarrow \infty} E(\exp\{i\beta Y(t)\}) = \lim_{t \rightarrow \infty} E(\exp\{i\beta(X(t) + \theta(t))\}) = E(\exp\{i\beta X\}).$$

**3. Main results.** Analogy of theorem 2 (and corollary 1) [7] takes place, which also can be proven using slight modification of results obtained in [2]:

**Claim 1.** If process  $\xi$  is supercritical,  $A_{\rho k}^l(u)$  and  $B_{\rho j k}^l$  are finite,  $i, j, k = \overline{1, n}$ , than random variables (r.v.)  $\frac{e^{-\rho t} \mu_1^i(t)}{K_1}, \dots, \frac{e^{-\rho t} \mu_r^i(t)}{K_r}, \frac{e^{-\rho t} \nu_{r+1}^i(t)}{K_{r+1}}, \dots, \frac{e^{-\rho t} \nu_n^i(t)}{K_n}$  converge as  $t \rightarrow \infty$  in square mean to the same limit  $\mu^i$ , where

$$K_j = \frac{v_\rho^j \int_0^\infty e^{-\rho t} (1 - G^i(t)) dt}{\rho \sum_{k,m=1}^r v_\rho^k u_\rho^m \int_0^\infty t e^{-\rho t} a_m^i(t) dG^k(t)}$$

$$K_l = \frac{\sum_{k=1}^r v_\rho^k \int_0^\infty e^{-\rho t} \int_0^t a_l^i(u) dG^i(u) dt}{\sum_{k,m=1}^r v_\rho^k u_\rho^m \int_0^\infty t e^{-\rho t} a_m^i(t) dG^k(t)}, \quad j = \overline{1, r}, \quad l = \overline{r+1, n}$$

Furthermore, Laplace transform  $\phi^l(s)$  of limit r.v.  $\tilde{\mu}^l$  satisfies equation

$$\phi^l(s) = \int_0^\infty h^l(u, \phi(se^{-\rho u})) dG^l(u),$$

with initial condition  $\phi(0) = 1, \phi'(0) = -1$ , where  $\phi(s) = (\phi^1(s), \dots, \phi^n(s))$ .

**Corollary 1.** *If conditions of theorem 1 are satisfied, than distributions*

$$P^i \left( \frac{\sum_{k=r+1}^n \nu_l^i(t)}{\sum_{k=1}^r \mu_k^i(t)} \leq x \mid \sum_{k=1}^r \mu_k^i(t) > 0 \right), i = \overline{1, r}$$

converge to the degenerate distribution, localized at the point

$$\frac{\sum_{l=r+1}^n \sum_{k=1}^r v_\rho^k \int_0^\infty e^{-\rho t} \int_0^t a_l^i(u) dG^i(u) dt}{\rho \sum_{l=1}^r v_\rho^l \int_0^\infty e^{-\rho t} (1 - G^i(t)) dt}.$$

In order to prove next theorem we will compare processes  $\nu_l^i(t)$  with processes  $N^i(t)$  — total number of particles born by the moment of time  $t$ , if at the moment  $t = 0$  there existed one particle of type  $T_i$ .

Let  $N_j^i(t)$  denote number of particles of type  $T_j$ , born by  $t$ , then  $N^i(t) = \sum_{j=1}^r N_j^i(t)$ .

It is known [8], that

$$E \left( \exp \left\{ i \sum_{j=1}^r \beta^j N_j^i(t) / v^j t^2 \right\} \mid \sum_{j=1}^r \mu_j^i(t) \right) \xrightarrow{t \rightarrow +\infty} \left( 2B \sum_{j=1}^r \beta^j \right)^{1/2} / M_a \left( sh \left( \left( 2B \sum_{j=1}^r \beta^j \right)^{1/2} / M_a \right) \right).$$

Then by letting  $\beta^j = v^j \beta$ , and since  $\sum_{j=1}^r v^j = 1$ , we get

$$E \left( \exp \left\{ i \beta N^i(t) / t^2 \right\} \mid \sum_{j=1}^r \mu_j^i(t) \right) \xrightarrow{t \rightarrow +\infty} (2B\beta)^{1/2} / M_a \left( sh \left( (2B\beta)^{1/2} / M_a \right) \right). \tag{6}$$

Also from [9], pp. 464–465, we can establish asymptotic behaviour of moments  $E^i(N_i(t))$  and  $E^i(N_i^2(t))$ :

$$E^i(N_i(t)) \sim \frac{u^i t}{M_a}, E^i(N_i^2(t)) \sim B \frac{u^i t^3}{3(M_a)^3}.$$

**Theorem 2.** *If the following conditions are satisfied:*

i) integrals  $M^j, M_a^{jk}, B_{jk}^i$  are finite;

- ii)  $\int_0^\infty \int_t^\infty a_l^m(u) dG^m(u) dt = k_l^m < +\infty$ ;  
 iii)  $\lim_{t \rightarrow \infty} t^2 \int_t^\infty a_m^k(u) dG^m(u) < +\infty$ ,  $\lim_{t \rightarrow \infty} t^2 (1 - G^l(t)) < +\infty$ ,  
 $m, k, j = \overline{1, r}, l = r + 1, n$ , then

$$E^i \left( \exp \left\{ i\beta \mu_i^i(t)/t^2 \right\} \mid \sum_{j=1}^r \mu_j^i(t) > 0 \right) \xrightarrow{t \rightarrow +\infty} \\ \xrightarrow{t \rightarrow +\infty} (2BM_0\beta)^{1/2} / M_a \left( sh \left( (2BM_0\beta)^{1/2} / M_a \right) \right).$$

**Proof.** As has been mentioned in (5), moments  $A_l^j(t) \sim \frac{u^i M_l t}{M_a}$ ,  $B_l^i(t) \sim \frac{u^i (M_l)^2 t^3}{3(M_a)^3}$ ,  $l = \overline{1, r}$ .

Now we will obtain asymptotic behavior of moments  $E^i(N^i(t)\nu_i^i(t))$ .

Let

$$F^i(t, z, s) = E^i \left( e^{zN^i(t)} s^{\bar{\mu}^i(t)} \right), z \leq 0, i = \overline{1, r},$$

$$F(t, z, s) = (F^1(t, z, s), \dots, F^r(t, z, s)).$$

Let also

$$D^i(t) = \frac{\partial F^i(t, z, s)}{\partial z} \Big|_{s=1, z=0}, D_l^i(t) = \frac{\partial^2 F^i(t, z, s)}{\partial z \partial s_l} \Big|_{s=1, z=0}, l = \overline{r+1, n}.$$

Similarly to [5], we can derive formula

$$F^i(t, z, s) = e^z \left( s_i (1 - G^i(t)) + \int_0^t h^i(u, F(t-u, z, s)) dG^i(u) \right). \quad (7)$$

Differentiating both sides of (7) first with respect to  $z$  at point 0, than with respect to  $s_0$  at the point  $s = \underbrace{(1, \dots, 1)}_n$  we get

$$D_l^i(t) = \int_0^t a_l^i(u) dG^i(u) + \sum_{k=1}^r \int_0^t A_l^k(t-u) a_k^i(u) dG^i(u) + \\ + \sum_{j,k=1}^r \int_0^t A_l^k(t-u) D^i(t-u) b_{jk}^i(u) dG^i(u) + \sum_{k=1}^r \int_0^t D_l^i(t-u) a_k^i(u) dG^i(u).$$

Using the results we obtained above, we see that

$$\int_0^t a_l^i(u) dG^i(u) + \sum_{k=1}^r \int_0^t A_l^k(t-u) a_k^i(u) dG^i(u) = o(t^2), \\ \sum_{j,k=1}^r \int_0^t A_l^k(t-u) D^i(t-u) b_{jk}^i(u) dG^i(u) \sim \sum_{j,k=1}^r B_{jk}^i u^j u^k \frac{M_l t^2}{(M_a)^2}.$$

So again, using **Theorem 1**, we obtain that  $D_l^i(t) \sim B \frac{u^i M_l t^3}{3(M_a)^3}$ .

Conditions ii) and iii) allows us to claim ([10]), that  $P^i \left( \sum_{k=1}^r \mu_k^i(t) > 0 \right) \sim \frac{2M_a}{Bt}$  and so

$$\begin{aligned} E \left( \left( \frac{(M_a)^2 N^i(t)}{t^2} - \frac{(M_a)^2 \nu_l^i(t)}{M_l t^2} \right)^2 \mid \sum_{k=1}^r \mu_k^i(t) > 0 \right) &\sim \\ &\sim \frac{BM_a u^i / 3t - 2BM_a u^i / 3t + BM_a u^i / 3t}{2M_a / Bt} = 0. \end{aligned}$$

From here, (6) and from **Lemma 1** we get the result.

**Corollary 2.** *If conditions of theorem 1 are satisfied, than distributions*

$$P^i \left( \frac{\nu_l^i(t)}{N^i(t)} \leq x \mid \sum_{k=1}^r \mu_k^i(t) > 0 \right),$$

converge to the degenerate distribution, localized at the point  $\frac{1}{M_l}$ , where,  $i = \overline{1, r}$ ,  $l = \overline{r+1, n}$ .

Analogy of theorem 1, [7] also takes place.

**Claim 2.** *Process  $\xi$  is critical or subcritical if and only if random variables  $\mu_l^j(t)$  converge with probability 1 as  $t \rightarrow +\infty$  to finitely valued  $r$ . v.  $\nu_l^j$ ,  $j = \overline{1, r}$ ,  $l = \overline{r+1, n}$ .*

Furthermore, vector characteristic function of limit r.v.  $\nu_l = (\nu_l^1, \dots, \nu_l^r)$  —

$$E(e^{i\beta\nu_l}) = \left( E(e^{i\beta\nu_l^1}), \dots, E(e^{i\beta\nu_l^r}) \right),$$

satisfies integral equation:

$$E(e^{i\beta\nu_l}) = \left( \int_0^{+\infty} h^1(u, \tilde{E}(e^{i\beta\nu})) dG^1(u), \dots, \int_0^{+\infty} h^r(u, \tilde{E}(e^{i\beta\nu})) dG^n(u) \right),$$

where

$$\tilde{E}(e^{i\beta\nu_l}) = \left( E(e^{i\beta\nu_l^1}), \dots, E(e^{i\beta\nu_l^r}), s_{r+1}, \dots, s_n \right).$$

**Proof.** Proof is analogical to theorem 1 in [7]. It is only worth to note that if process is supercritical, then there exist  $i = \overline{1, r}$ , such that

$$0 < \lim_{t \rightarrow +\infty} P^i \left( \sum_{k=1}^r \mu_k^i(t) > 0 \right) = \lim_{t \rightarrow +\infty} Q^i(t) = q^i < 1,$$

and process  $N^i(t) \xrightarrow{t \rightarrow +\infty} +\infty$  (see chapter 6, [5]) with probability  $q^i$ , so the strong law of large numbers holds for  $\frac{\nu_l^i}{N^i}$ .

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**Лисецький Т. Б.** Про кількість емігрантів в розкладному гіллястому процесі з перетвореннями, залежними від віку.

Розкладний гіллястий процес можна розглядати як стохастичну модель популяції з  $N$  типами індивідумів, розділених на декілька підгруп  $G_1, G_2, \dots, G_n$ ,  $n \leq N$ , де кожна група населяє окремий острів. Індивідум з групи  $G_i$  може одразу після народження емігрувати на острів, населений групою з вищим індексом, або залишитись на своєму острові. В даній статті розглядається випадок з двома групами  $G_1$  та  $G_2$ . Кожна особа має випадкову тривалість життя, а розподіл її потомства залежить від її віку.

Ми досліджуємо асимптотичну поведінку процесів, які рахують кількість частинок, що емігрували, в залежності від критичності гіллястого підпроцесу, породженого групою  $G_1$ .

**Ключові слова:** гіллястий процес, стохастичний адитивний функціонал, критичний гіллястий процес, перонів корінь, моменти.

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